

Rapid computation of prices and deltas of nth to default swaps in the Li Model

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Summary

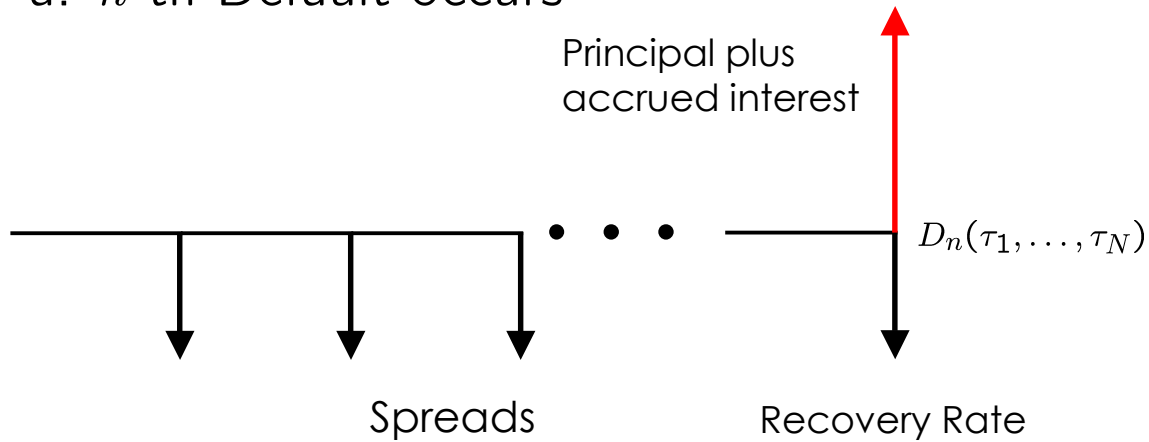
- Basic description of an nth to default swap
- Introduction to the Li model
- Solutions: Importance Sampling
- Parameter hedging and why computing sensitivities are difficult.
- Solutions: Likelihood & Pathwise Methods
- Results

Nth to default swaps: Product definition

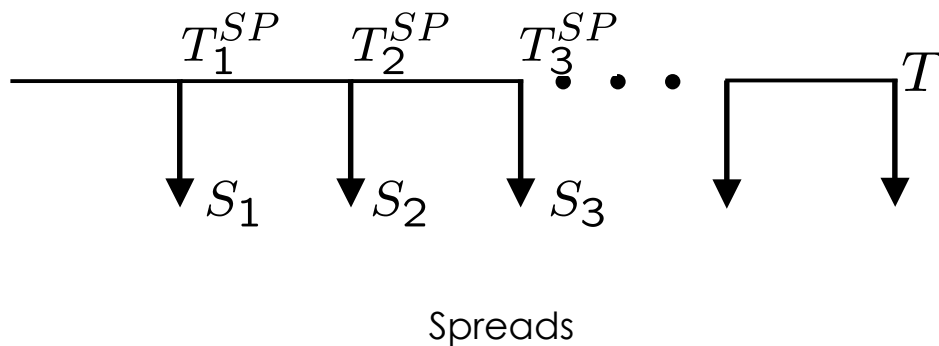
- In an n th default swap a regular fee is paid until n of a basket of N credits have defaulted, or the deal finishes.
- When the N th default occurs a payment of $1 - R$ is made to the fee payer.
 R = recovery rate of n th defaulting asset

Nth to default swaps: Product definition

a. n th Default occurs



b. n th Default does not occur



The Li Model

- Defaults are assumed to occur for individual assets according to a Poisson process with a deterministic intensity called the *hazard rate*.
- This means that default times are exponentially distributed.
- Li: Correlate these default times using a Gaussian copula

Some Definitions

- Consider some security A. We define the **default time**, τ_A , as the time from today until A defaults.
- We assume the defaults to occur as a Poisson process
- The intensity of this process, $h(t)$, is called the hazard rate.

The Pricing Algorithm: SetUp

Given a correlation matrix C we compute A such that

$$AA^T = C$$

Let $E(\tau, h)$ denote the cumulative exponential distribution function in τ given a fixed h :

$$E(\tau, h) = \mathbb{P}(t < \tau) = 1 - \exp\left(-\int_0^\tau h_j(t) dt\right).$$

$E^{-1}(u, h)$ denotes its inverse for fixed h .

The Pricing Algorithm

- Draw a vector of independent normals, \mathbf{z}
- Generate a set of correlated Gaussian deviates:

$$\mathbf{w} = A\mathbf{z}.$$

- Map to uniforms:

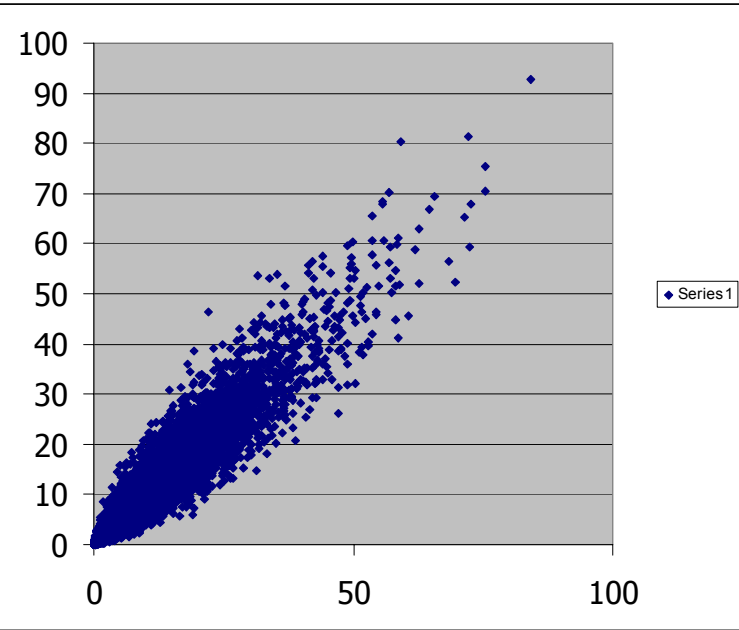
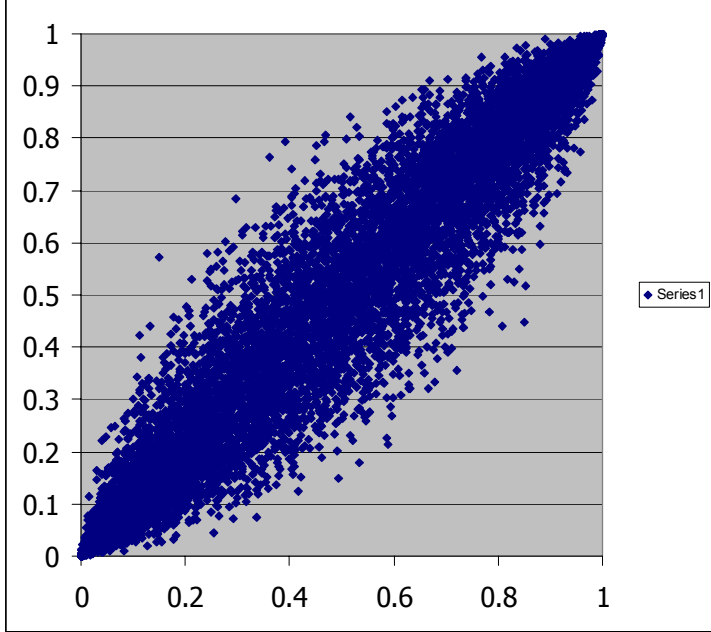
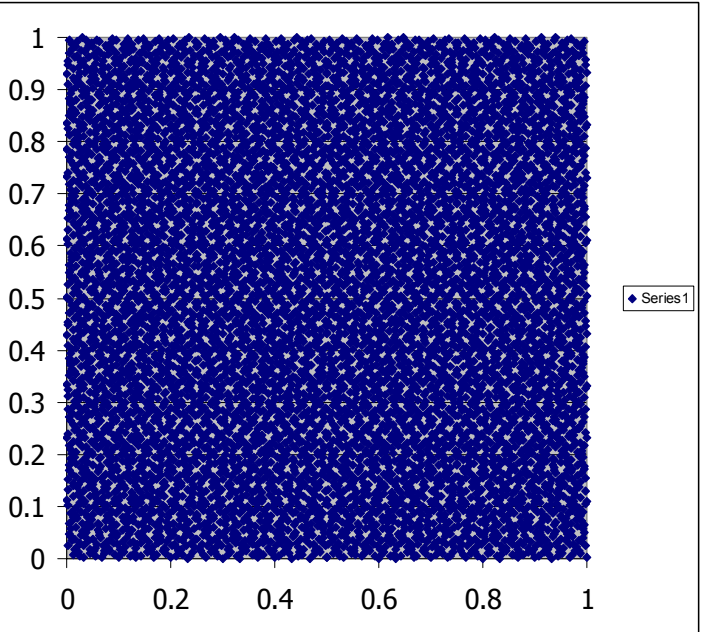
$$u_i = N(w_i)$$

- Map to default times:

$$\tau_i = E^{-1}(u_i, h)$$

- Compute the cash flow in this scenario; discount back.

$$V^F(\tau_1, \dots, \tau_N) = P(D_n(\tau_1, \dots, \tau_N)) [V_{\text{prot}} + (1 - r_n)H(T - D_n(\tau_1, \dots, \tau_N))].$$



Importance Sampling

Intuitively: want to sample more thoroughly in the regions where defaults occur.

Look at a k th to default swap:

- Product pays a constant amount unless k defaults occur.
- Restrict our attention to cases of k defaults.
- By subtracting the constant, we can assume value is zero unless k defaults occur.

Importance Sampling

- General Strategy: alter the probabilities of default such that we always get k defaults. Each path is then “important”; compute prices.
- We then reweight the different contributions according to our changes to the probability measure

Designing the importance density when $i = 1$

Make the i th asset default before T with probability:

$$\frac{1}{(n + 1) - i}$$

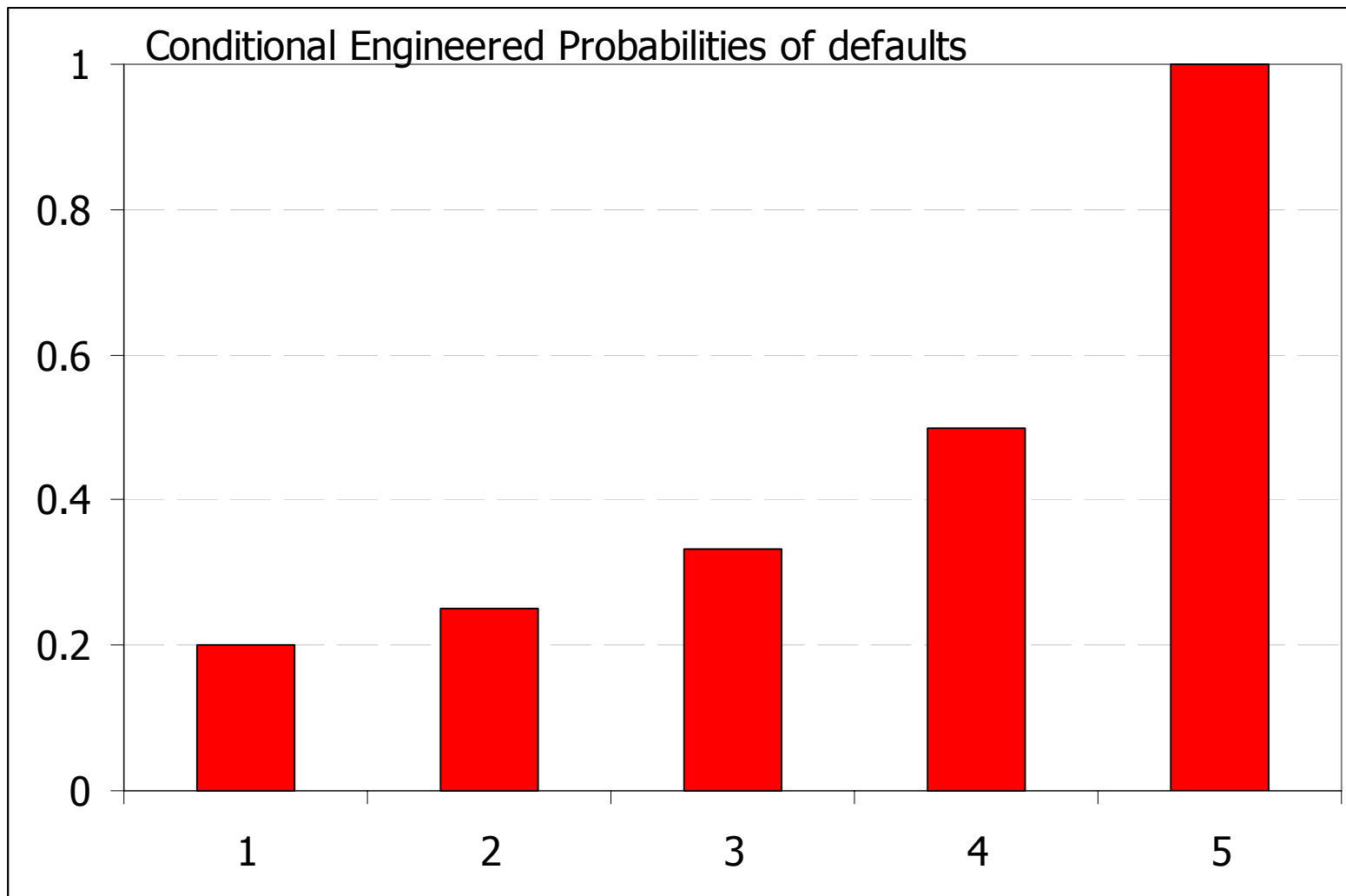
- Why? After i non defaults want all the remaining credits to have an equal chance of default

- Pick a uniform u_i . If:

$u_i < \frac{1}{n + 1 - i}$ map u_i to a region where asset i defaults.

$u_i > \frac{1}{n + 1 - i}$ map u_i a region where asset i doesn't default.

Designing the importance density when $i = 1$



Designing the importance density



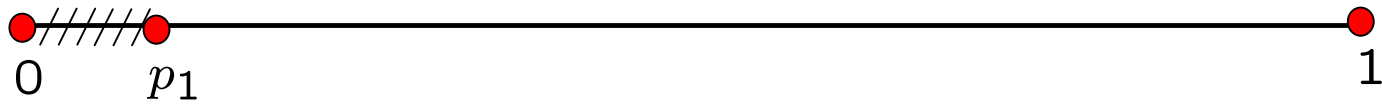
• Look at the original default region for asset i

$$\tau_i < T \quad \longrightarrow \quad w_i < x \quad \longrightarrow \quad \sum_j A_{ij} z_j < x$$

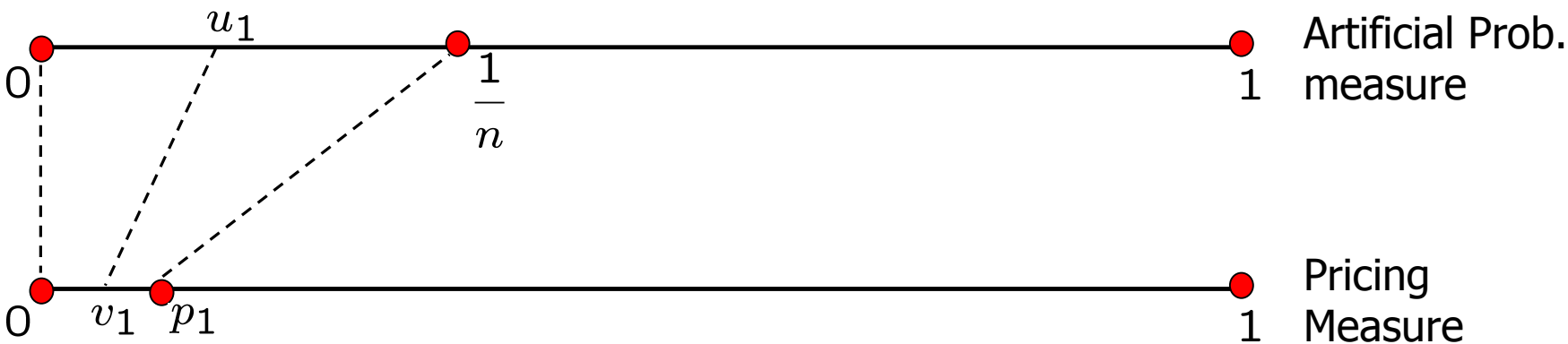
↑ Correlated Gaussian
 ↑ $N(0, 1)$

• For our first to default case: $a_{11} z_1 < x \implies z_1 < \frac{x_1}{a_{11}}$

• Translate to uniforms: $p_1 = N\left(\frac{x_1}{a_{11}}\right)$

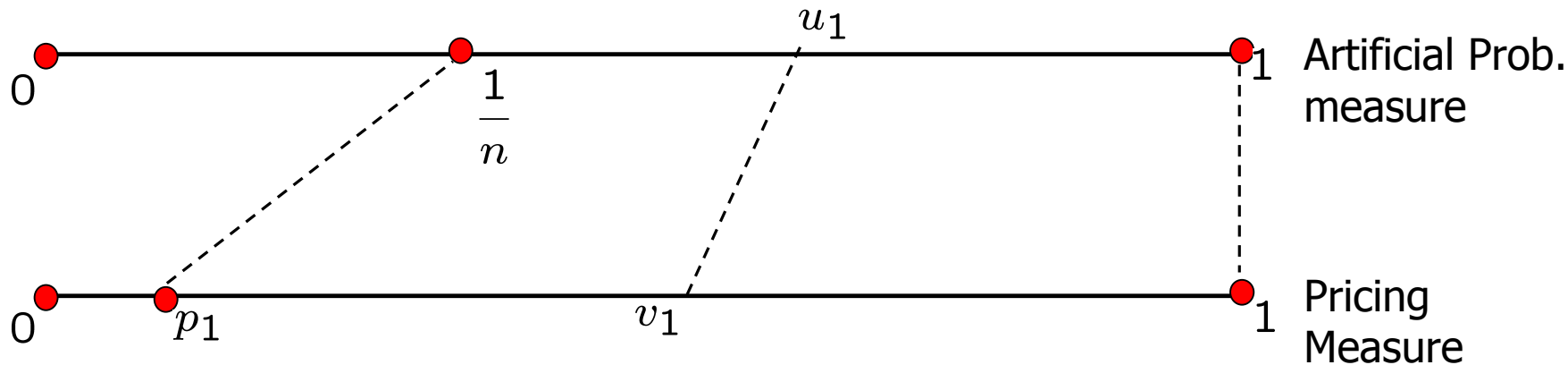


rst to Default occurs:



• u_1 maps to v_1 where: $\frac{v_1}{p_1} = u_1 n \implies v_1 = u_1 n p_1$

rst to Default doesn't occur:



• u_1 maps to v_1 where: $v_1 = p_1 + \frac{1 - p_1}{1 - \frac{1}{n}} \left(u_1 - \frac{1}{n} \right)$

We need to scale the contributions of these paths

First asset defaults: weight by np_1

Doesn't default: weight by $\frac{1 - p_1}{1 - \frac{1}{n}}$

Suppose that we have dealt with the first $(j-1)$ assets. The unmassaged default probability now depends on Z :

$$W_j < x_j \text{ if and only if } \sum_{i < j} a_{ij} Z_i + a_{jj} Z_j < x_j.$$

However, as A is lower triangular we have

$$p_j = \frac{x_j - \sum_{i < j} a_{ij} Z_j}{a_{jj}}$$

And repeat as before.

Computing Hazard Rate Sensitivities

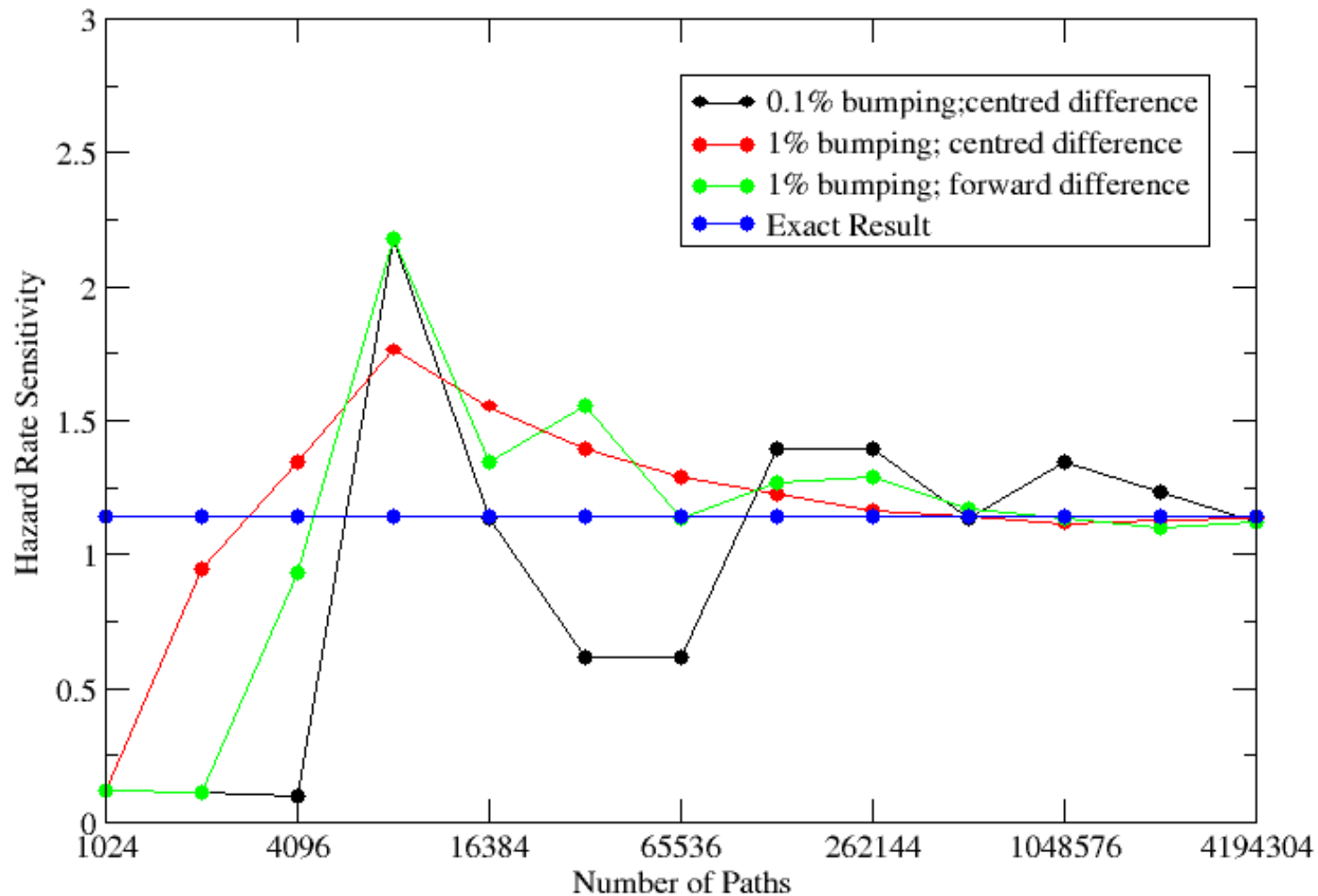
- We hedge against changes in the hazard rates of the individual assets using “vanilla” default swaps.
- Naïve methods for determining hazard rate sensitivities (finite differencing)

$$\Delta = \frac{P(h_i + \epsilon) - P(h_i)}{\epsilon} \quad \text{or} \quad \Delta = \frac{P(h_i + \epsilon) - P(h_i - \epsilon)}{2\epsilon}$$

have severe limitations due to their (very) slow rate of convergence.

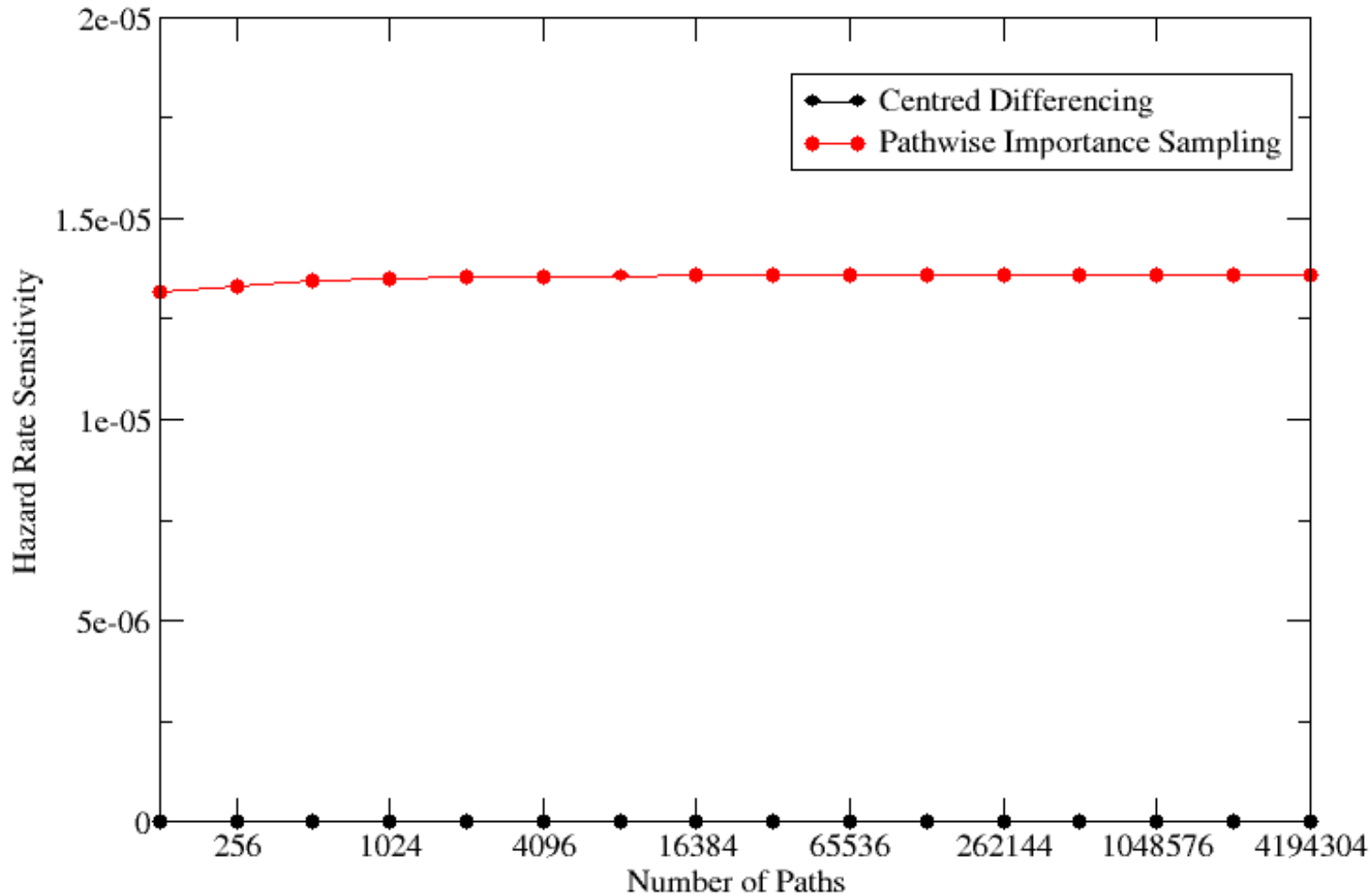
Computing Hazard Rate Sensitivities

First to default, 4 credits, 2 year deal **Not a stress case !**



Computing Hazard Rate Sensitivities

Fourth to default, 4 credits, 0.15 year deal



Why is Bumping problematic ?

- Very few paths will give multiple defaults a short time (e.g 0.15 years). If obligors are uncorrelated,

$$\text{Prob } n \text{ defaults} = (hT)^n$$

We therefore need lots of paths, even for pricing.

- When we compute sensitivities, bump one hazard rate. Very small change in the number of paths which now have n defaults compared to previously.

Why is Bumping problematic ?

A CDS is similar to a barrier option, pay-out jumps according to whether Nth default is before or after deal maturity.

Value CDS =

$$\int P(D_n(\tau_1, \dots, \tau_N)) [(1-r_n)H(T-D_n(\tau_1, \dots, \tau_N))\psi(\tau_1, \dots, \tau_N)] d\tau_1 \dots d\tau_N.$$

When we differentiate the payoff w.r.t the hazard rates we get a δ function.

Sampling this by Monte Carlo is very hard.

Parameter Sensitivities Using Monte Carlo

Well-known techniques for computing Greeks by Monte Carlo include:

- **Likelihood ratio**: differentiate the probability density function analytically, inside the integral.
- The **Pathwise Method**: differentiate the Payoff.
Generally believed not to apply to discontinuous payoffs
– we show that it does apply.
Broddie-Glasserman
- **Malliavin calculus**: differentiation w.r.t. the underlying Brownian motion; not applicable here.

The Likelihood Ratio Method

Value of the option:

$$V = \mathbb{E}^{\mathbb{Q}}[F(S_T)] = \int F(S)\psi(S, \theta) dS$$

We can write the sensitivity w.r.t θ :

$$\frac{\partial V}{\partial \theta} = \int F(S) \frac{\partial}{\partial \theta} \psi(S, \theta) dS$$

No longer integrating against our Monte Carlo density!
However, we can reintroduce it:

The Likelihood Ratio Method

$$\begin{aligned}\frac{\partial V}{\partial \theta} &= \int F(S) \frac{\partial \psi(S, \theta)}{\partial \theta} \frac{1}{\psi(S, \theta)} \psi(S, \theta) dS \\ &= \int F(S) \frac{\partial}{\partial \theta} \log \psi(S, \theta) \psi(S, \theta) dS\end{aligned}$$

∴ To compute sensitivity we reweight the payoff with:

$$\frac{\partial}{\partial \theta} \log \psi(S, \theta)$$

The Pathwise Method

- The delta of an option with payoff $F(S_T)$ is:

$$\Delta = \frac{\partial V}{\partial S_0} = e^{-rT} \int F(S_T) \frac{\partial}{\partial S_0} \psi(S_T, S_0, \dots) dS_T$$

- For the case of a lognormal evolution we can show:

$$\Delta = \frac{\partial \psi}{\partial S_0} = -\frac{\partial}{\partial S_T} \left(\frac{S_T}{S_0} \psi \right)$$

- Integrating by parts and eliminating the boundary term:

$$\Delta = e^{-rT} \int \frac{\partial F(S_T)}{\partial S_T} \frac{S_T}{S_0} \psi(S_T, S_0, \dots) dS_T$$

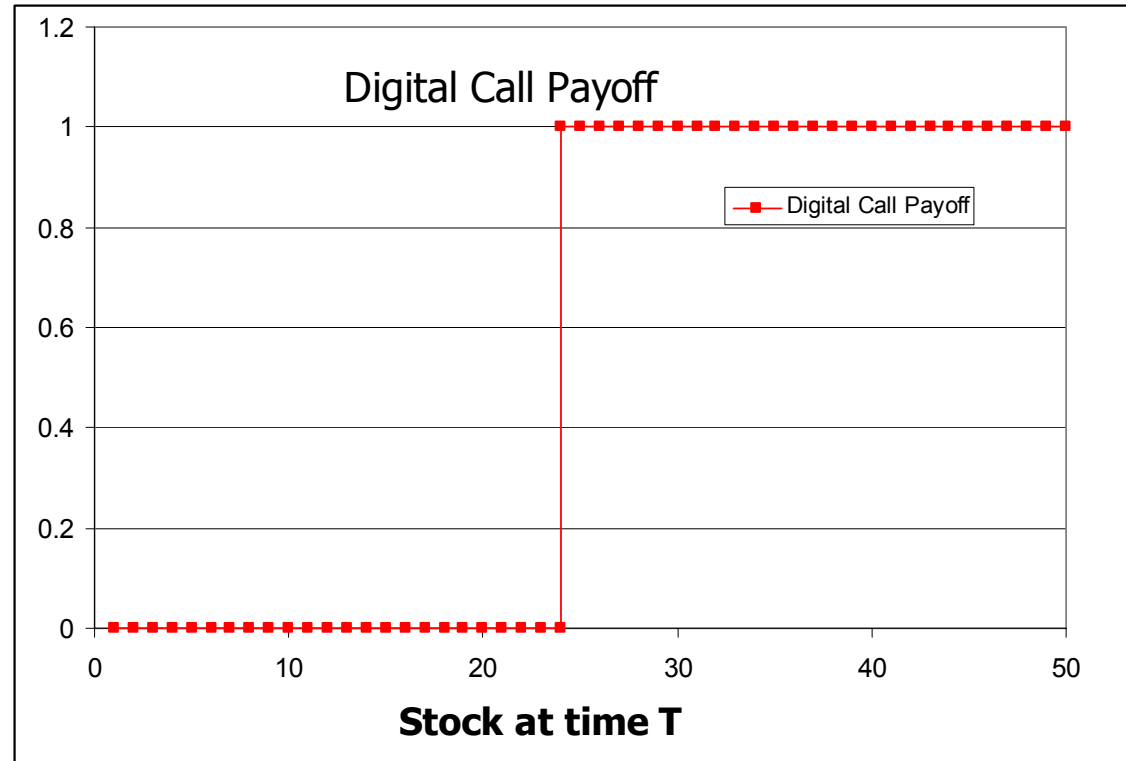
The Pathwise Method

- We are now differentiating the *payoff*!

- Suppose we have a digital option:

$$f(S_T) = H(S_T - K)$$

- Differentiate and we get a δ function



The Likelihood Ratio Method for nth Default Swaps

- Value of the CDS:

$$\int P(D_n)(1-r_n)H(T-D_n)\psi(\tau_1, \dots, \tau_N)d\tau_1 \dots d\tau_N.$$

- Differentiate w.r.t. i th hazard rate :

$$\frac{\partial V}{\partial h_i} = \int_0^T P(D_n)(1-r_n)H(T-D_n)\frac{\partial \psi(\tau_1, \dots, \tau_N)}{\partial h_i}d\tau_1 \dots d\tau_N.$$

- Applying Broadie/Glasserman's trick:

$$\frac{\partial V}{\partial h_i} = \int_0^T P(D_n)(1-r_n)H(T-D_n)\frac{\partial \log \psi(\tau_1, \dots, \tau_N)}{\partial h_i}\psi(\tau_1, \dots, \tau_N)d\tau_1 \dots d\tau_N.$$

The Likelihood Ratio Method for nth Default Swaps

- The calculation is straightforward for Gaussian copula and flat hazard rates:

$$\frac{\partial \log \psi(\tau_1, \dots, \tau_n)}{\partial h_i} = -(\rho^{-1} - \mathbf{1})_{ij} \eta_j \frac{\partial \eta_i}{\partial u_i} \frac{\partial u_i}{\partial h_i} + \frac{1}{h_i} - \tau_i$$

where ρ is the correlation matrix and

$$\eta_i = \phi^{-1}(u_i) \quad \frac{\partial \eta_i}{\partial u_i} = \sqrt{2\pi} e^{-\frac{1}{2} \phi^{-1}(u_i)^2}$$

The Pathwise Method for nth Default Swaps

• We differentiate the *discounted pay-off* w.r.t h_j (ignore the spreads for the moment):

$$F(\tau_1, \dots, \tau_N) = P(D_N(\tau_1, \dots, \tau_N))[(1 - r_n)H(T - D_n(\tau_1, \dots, \tau_N))]$$

$$\frac{\partial F}{\partial h_j} = \frac{\partial F}{\partial \tau_j} \frac{\partial \tau_j}{\partial h_j}$$

where if the j th asset is the n th to default

$$\begin{aligned} \frac{\partial F}{\partial \tau_j} = & \frac{\partial P}{\partial t}(\tau_j)[H(T - \tau_j)(1 - r_N)] \\ & - P(\tau_j)[\delta(\tau_j - T)(1 - r_n) + H(\tau_j - T) \frac{\partial}{\partial t}(1 - r_n)|_{t=\tau_j}] \end{aligned}$$

And zero otherwise.

The Pathwise Method for n th Default Swaps

The important terms are the second and third terms.

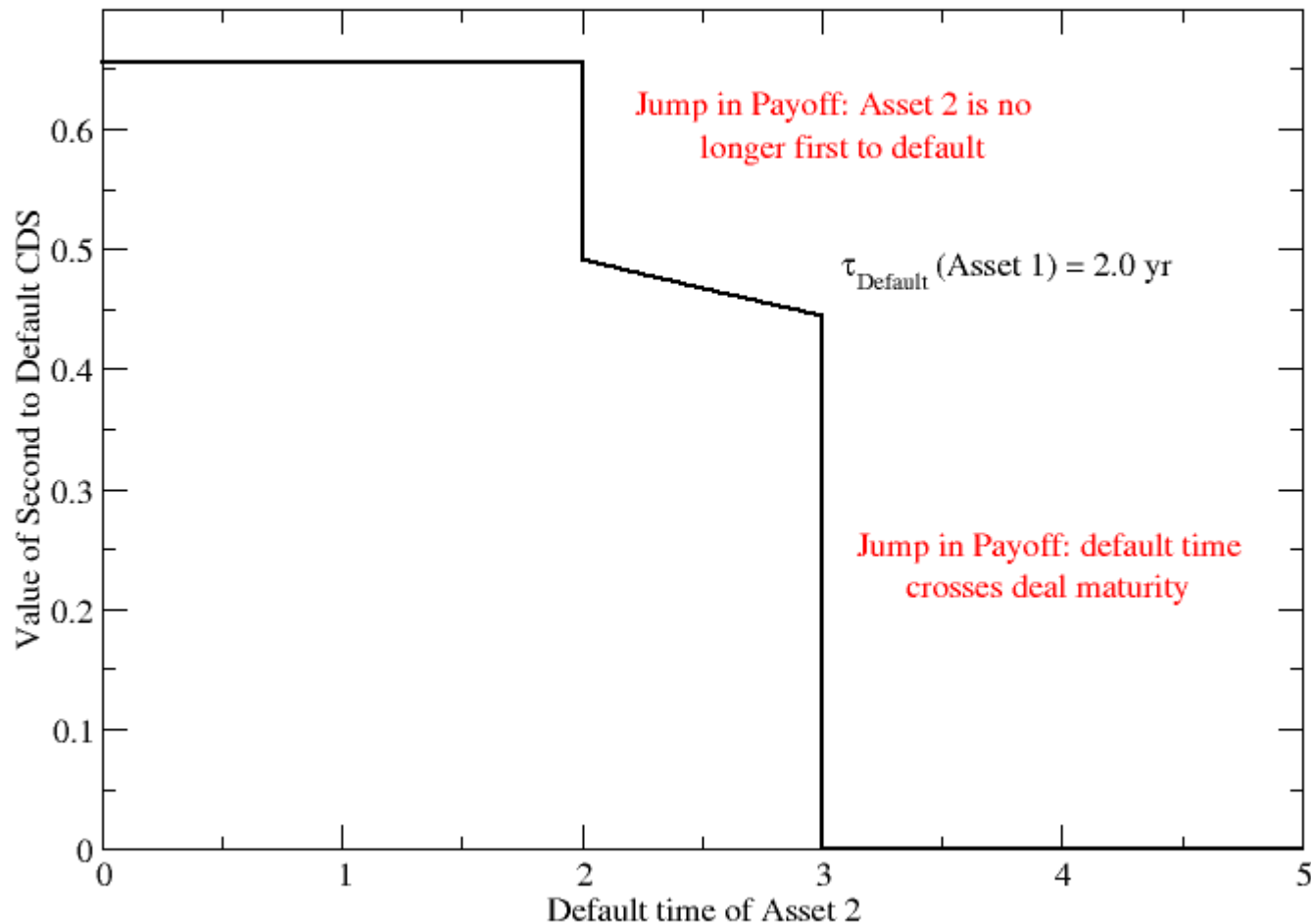
They correspond to:

a. default time of j th asset crosses final maturity of the product.

b. Upon bumping the j th hazard rate we alter which asset is the n th to default

Both result in a jump in value and hence a Delta function in the derivative.

The Pathwise Method for nth Default Swaps



- When differentiated these jumps in the payoff give rise to delta functions !

The Pathwise Method for nth Default Swaps

The delta functions make a bumped Monte Carlo converge very slowly. However, we can integrate these *analytically* to obtain

$$-P(T) \frac{\partial E^{-1}}{\partial h_j} \int \psi(\tau_1, \dots, \tau_{j-1}, T, \tau_{j+1}, \dots, \tau_N) d\tau_1 \dots d\tau_{j-1} d\tau_{j+1} \dots d\tau_n.$$

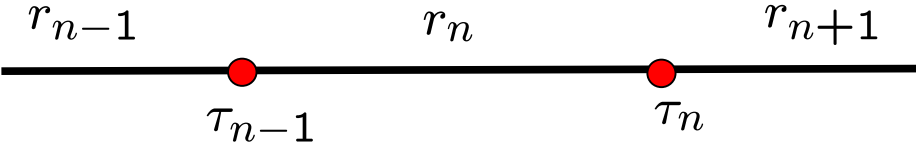
As before we simply reintroduce it, the second term is now

$$\int \frac{(I\psi(\tau_1, \dots, \tau_{j-1}, T, \tau_{j+1}, \dots, \tau_N))}{\psi_{n-1}(\tau_1, \dots, \tau_{j-1}, \tau_{j+1}, \dots, \tau_N)} \psi_{n-1}(\tau_1, \dots, \tau_{j-1}, \tau_{j+1}, \dots, \tau_N) d\tau_1 \dots d\tau_{j-1} d\tau_{j+1} \dots d\tau_n,$$

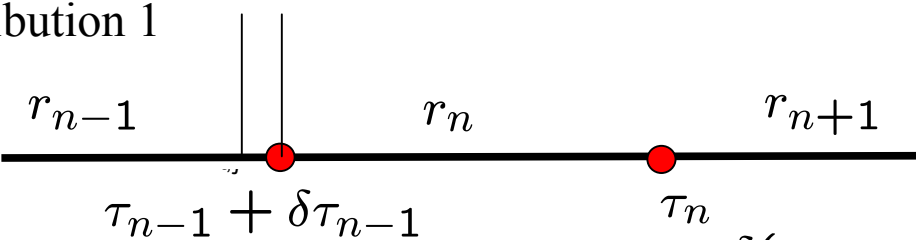
where $I = 1$ if t_j is the nth default time and zero otherwise.

Delta contributions from recovery rates

Two possible contributions: after sorting j th bond becomes $(n-1)$ th or n th default.

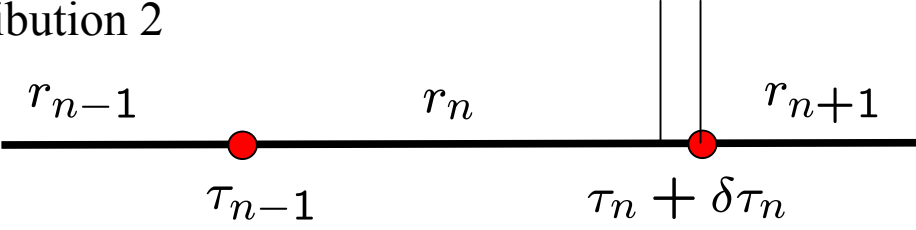


Contribution 1



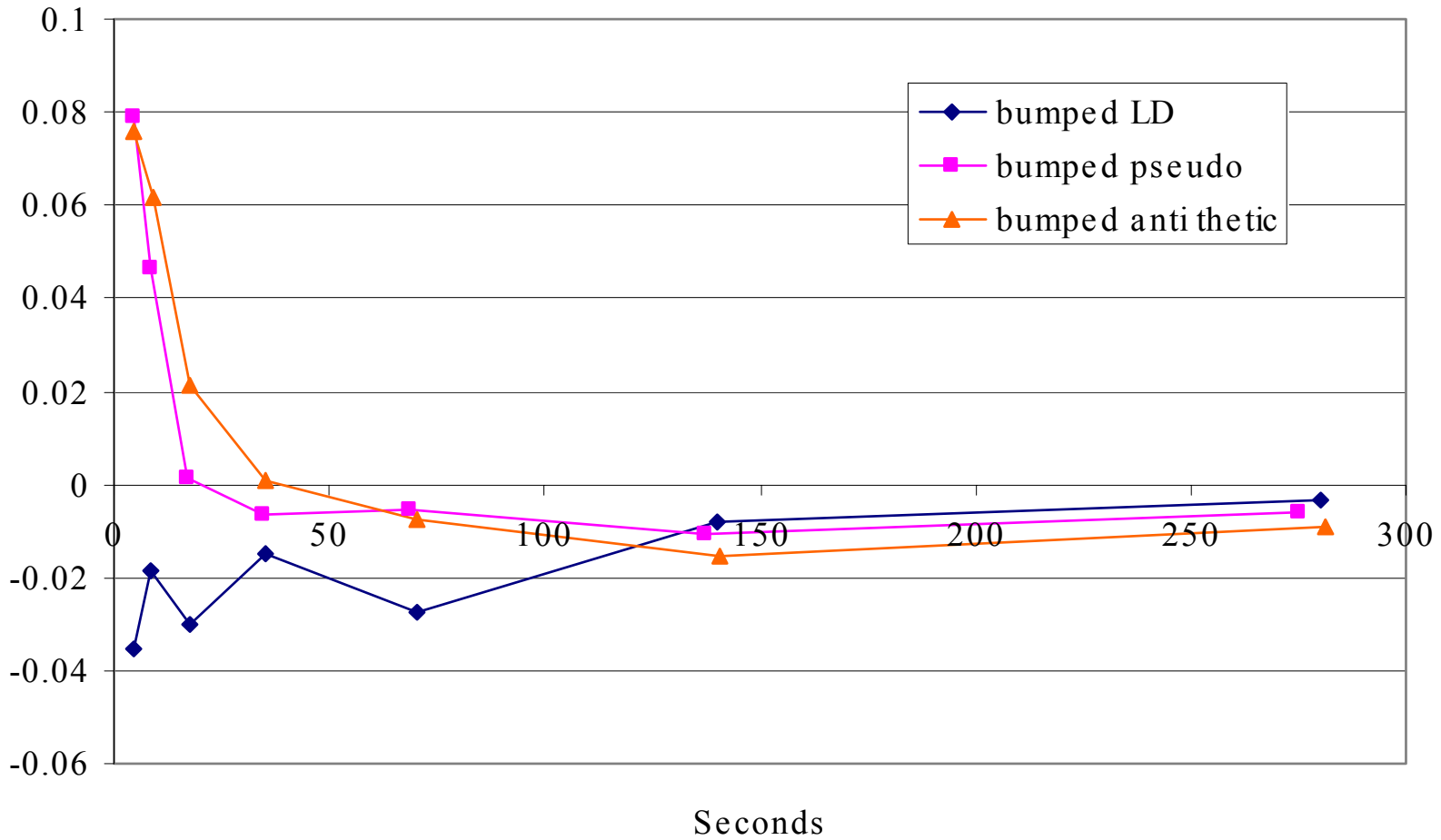
$$\delta(\tau_{n-1} - T)[((1 - r_n) - (1 - r_{n-1}))P(T)]$$

Contribution 2

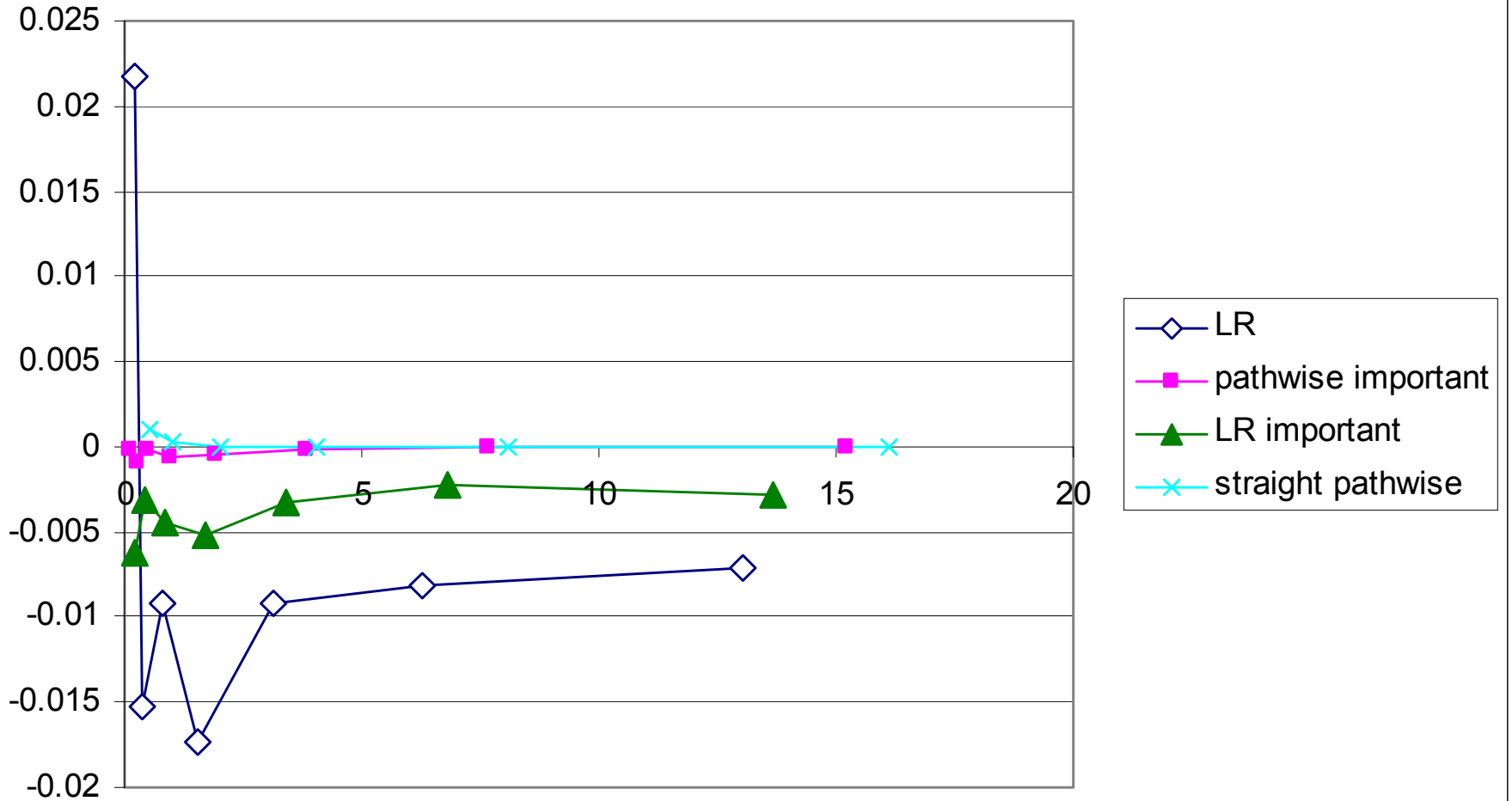


$$\delta(\tau_n - T)[((1 - r_{n+1}) - (1 - r_n))P(T)]$$

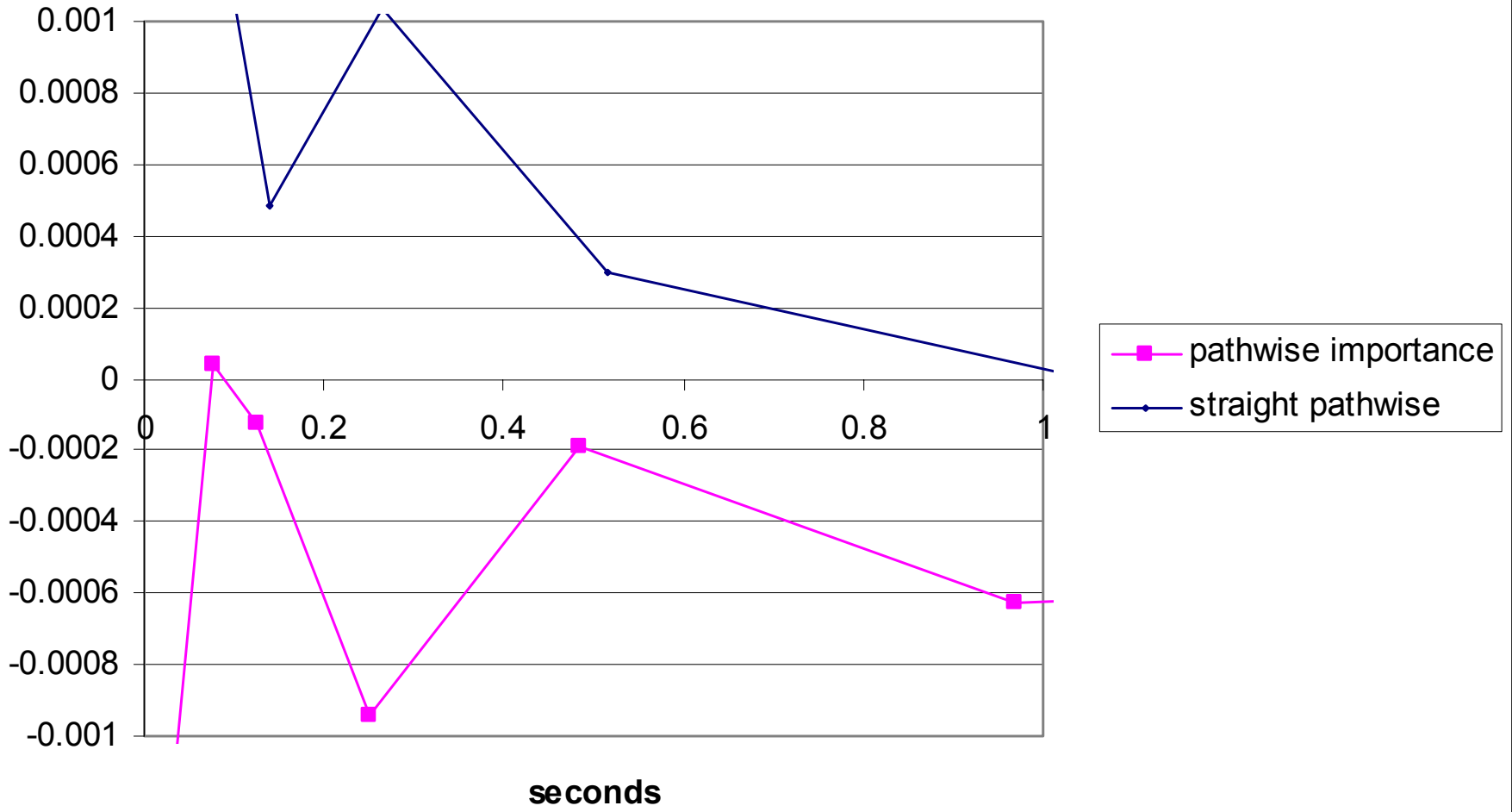
Error in convergence of first to default on 4 assets five year deal, 2 percent hazard rates, value of delta 2.015



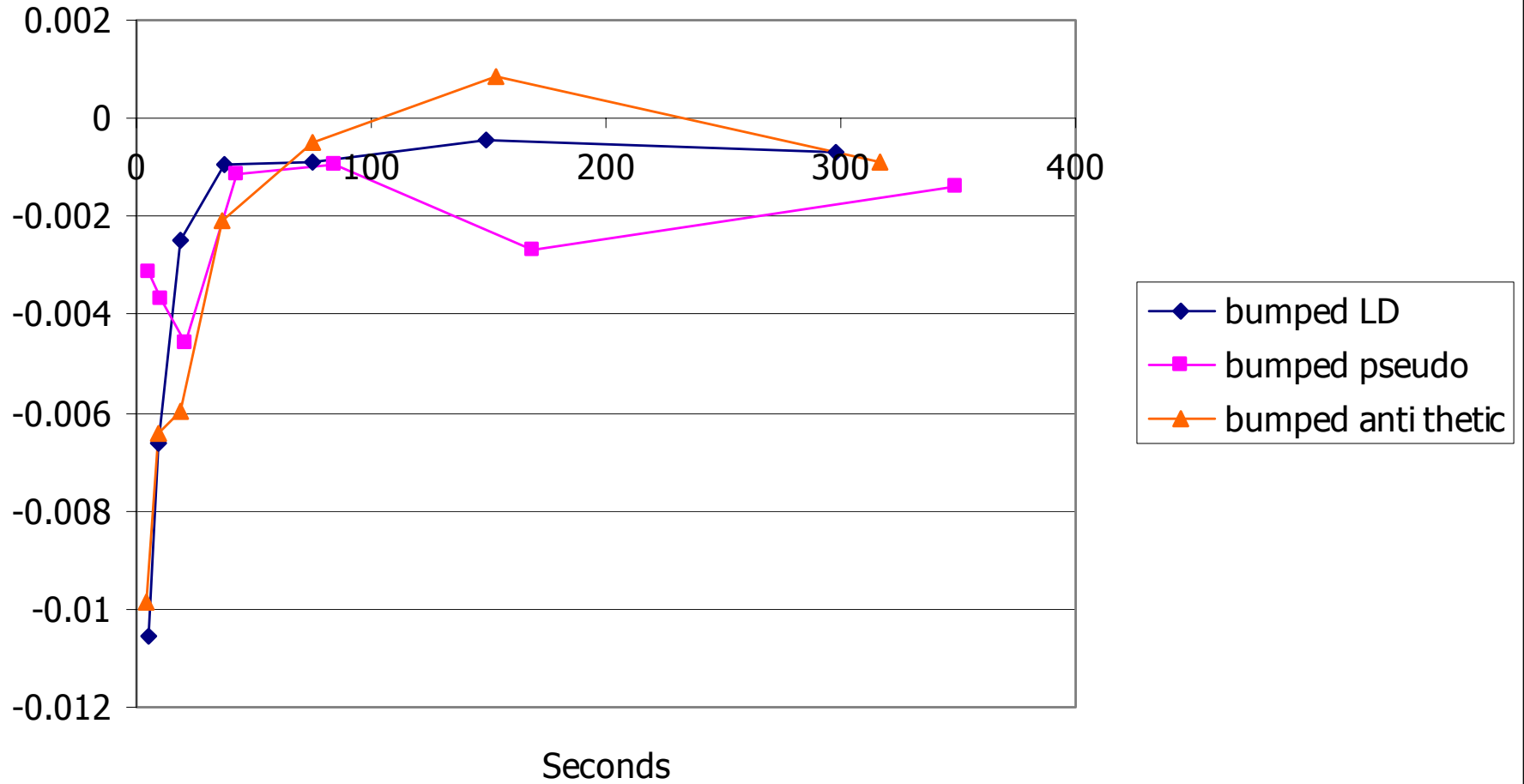
Error in convergence of first to default on 4 assets five year deal, 2 percent hazard rates, value of delta 2.015



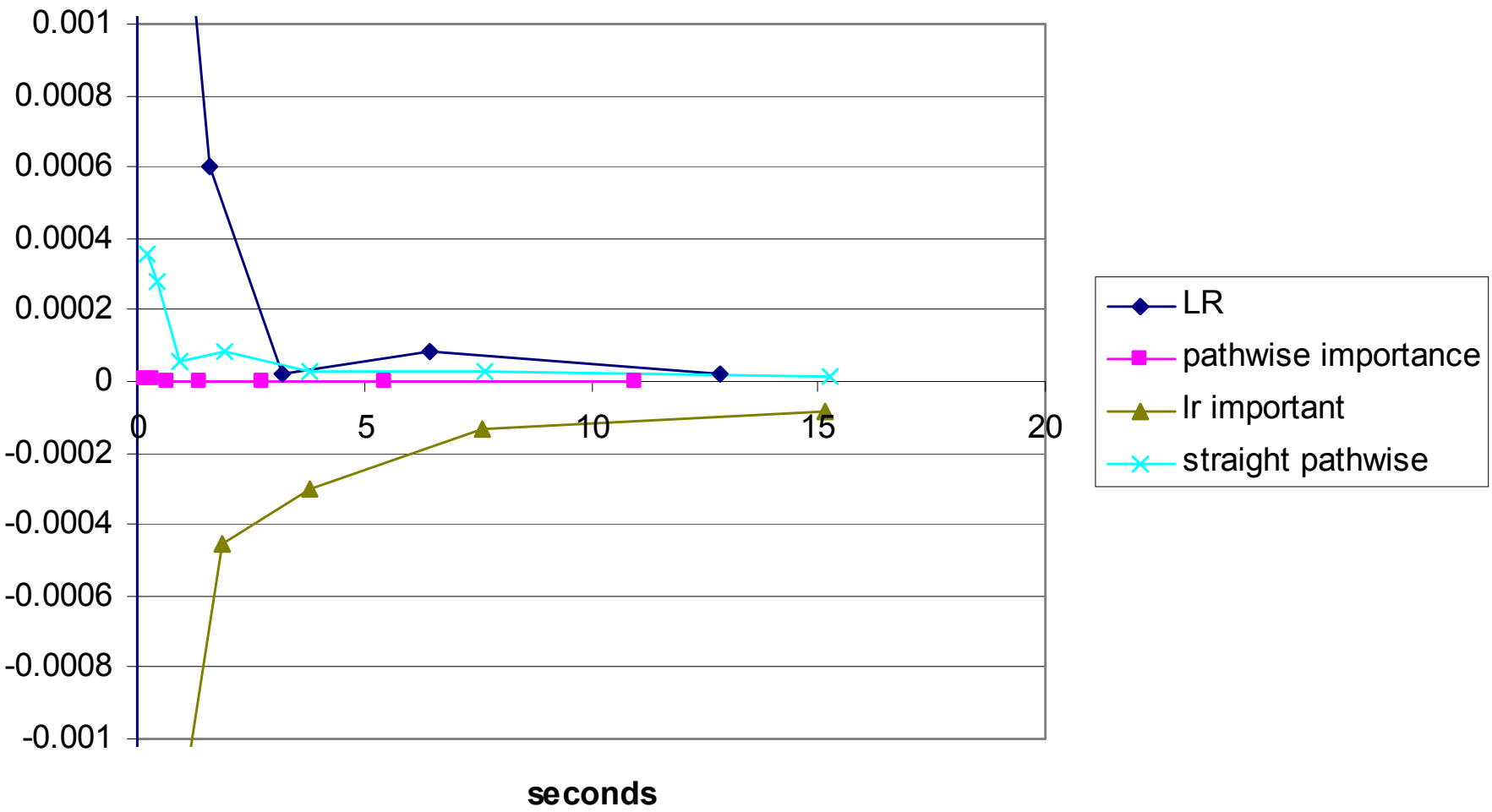
Error in convergence for first to default on 4 assets, five year deal, 2 percent hazard rates, value of delta 2.015



Error in convergence of fourth to default on 4 assets five year deal, 2 percent hazard rates, value of delta 0.01557



Error in convergence for fourth to default on 4 assets, five year deal, 2 percent hazard rates, value of delta 0.01557



General Results

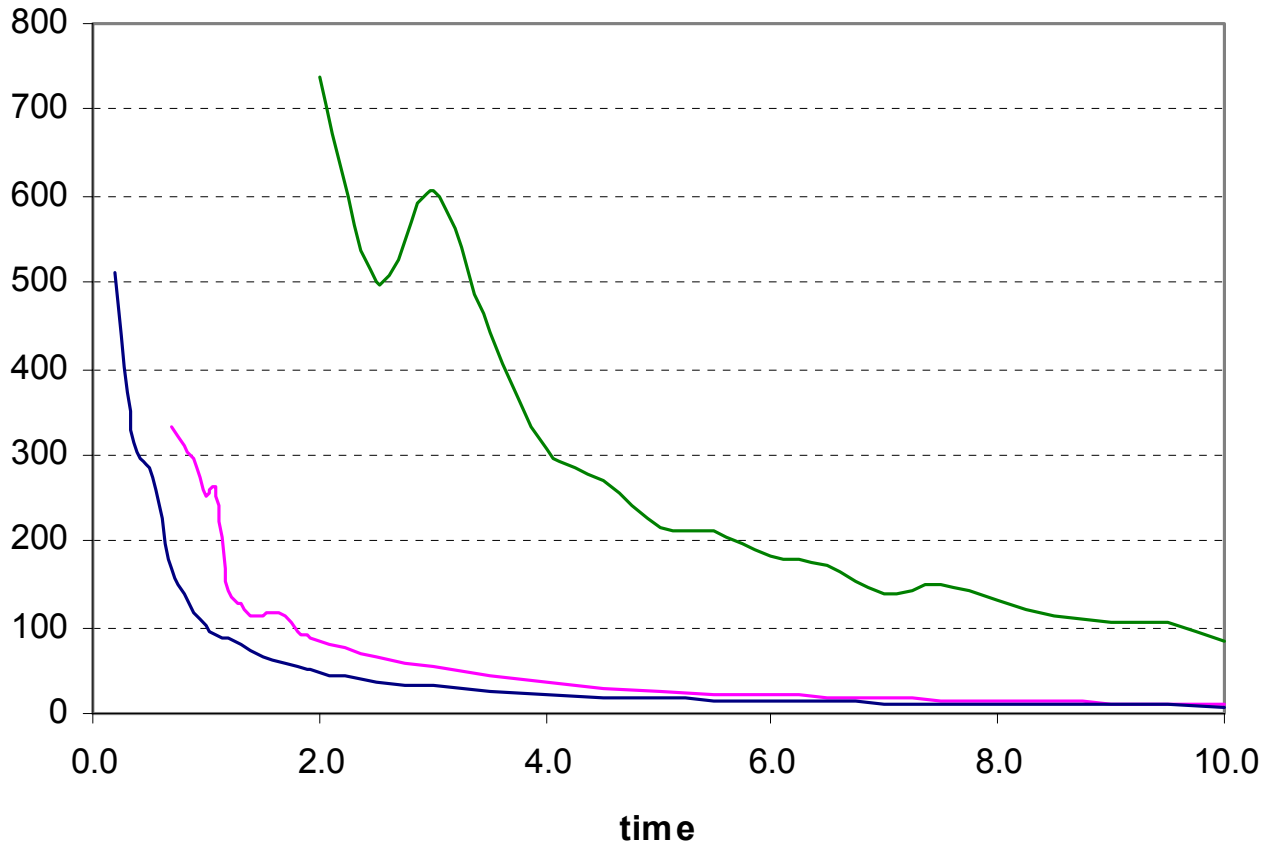
If we run a Monte Carlo simulation for n paths then the standard error is

$$\frac{\sigma}{\sqrt{n}}$$

where σ is the standard deviation.

In the following, we therefore plot the standard deviation of the result as a fraction of the result.

standard deviation of delta as a fraction of delta with varying maturity for fourth to default with four assets with varying recovery rates (protection leg only)



- likelihood ratio
- bumped
- pathwise

standard deviation of delta as a fraction of delta with varying maturity for fourth to default with four assets with varying recovery rates (protection leg only)

